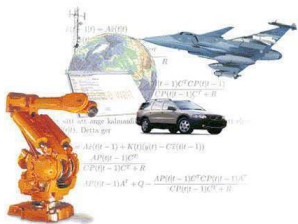


# Nonlinear system identification enabled via sequential Monte Carlo



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Joint work with (alphabetical order): **Michael I. Jordan** (UC Berkeley), **Fredrik Lindsten** (Linköping University), **Lennart Ljung** (Linköping University), **Brett Ninness** (University of Newcastle, Australia) and **Adrian Wills** (MRA, Newcastle, Australia).



A state space model (SSM) consists of a Markov process  $\{x_t\}_{t \geq 1}$  and a measurement process  $\{y_t\}_{t \geq 1}$ , related according to

$$\begin{aligned}x_{t+1} \mid x_t &\sim f_{\theta,t}(x_{t+1} \mid x_t), \\y_t \mid x_t &\sim g_{\theta,t}(y_t \mid x_t), \\x_1 &\sim \mu_{\theta}(x_1), \\(\theta &\sim p(\theta)).\end{aligned}$$

We observe

$$y_{1:T} \triangleq \{y_1, \dots, y_T\},$$

(leaving the latent variables  $x_{1:T}$  unobserved).

**Identification problem:** Find  $\theta$  based on  $y_{1:T}$ .



Alternate between updating  $\theta$  and updating  $x_{1:T}$ .

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### Frequentists:

- Find  $\hat{\theta}_{\text{ML}} = \arg \max_{\theta} p_{\theta}(y_{1:T})$ .
- Use e.g. the expectation maximization (EM) algorithm.

### Bayesians:

- Find  $p(\theta | y_{1:T})$ .
- Use e.g. Gibbs sampling.



1. **Maximum Likelihood (ML) identification**
    - Problem formulation
    - Solution using EM and a particle smoother
  2. **Bayesian identification**
    - Problem formulation
    - Gibbs sampling
  3. Sequential Monte Carlo (SMC), the particle filter
  4. Particle Gibbs with ancestor sampling (PG-AS)
- 

The sequential Monte Carlo samplers are fundamental to **both** the ML and the Bayesian approaches.



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**Identification problem:** Find  $\theta$  based on  $y_{1:T}$ .

ML amounts to solving,

$$\hat{\theta}^{\text{ML}} = \arg \max_{\theta} \log p_{\theta}(y_{1:T})$$

where the log-likelihood function is given by

$$\log p_{\theta}(y_{1:T}) = \sum_{t=1}^T \log p_{\theta}(y_t \mid y_{1:t-1})$$



The **EM** algorithm computes ML estimates of unknown parameters in probabilistic models involving latent variables.

The latent variables in an SSM are given by the states,

$$\{x_1, \dots, x_T\}.$$

**Strategy:** Use the *structure* inherent in the SSM to separate the original problem into *two closely linked subproblems*, each of which is hopefully in some sense more tractable than the original problem.



**Algorithm 1** EM for identifying nonlinear dynamical systems

1. **Initialise:** Set  $i = 1$  and choose an initial  $\theta^1$ .
2. **While** not converged **do:**

(a) **Expectation (E) step:** Compute

$$\begin{aligned} Q(\theta, \theta^i) &= E_{\theta^i} [\log p_{\theta}(x_{1:T}, y_{1:T}) \mid y_{1:T}] \\ &= \int \log p_{\theta}(x_{1:T}, y_{1:T}) \underbrace{p_{\theta^i}(x_{1:T} \mid y_{1:T})}_{\text{}} dx_{1:T} \end{aligned}$$

using **PS** (forward filter/**backward simulation**, FFBS).

- (b) **Maximization (M) step:** Compute  $\theta^{i+1} = \arg \max_{\theta \in \Theta} Q(\theta, \theta^i)$
- (c)  $i \leftarrow i + 1$

Thomas B. Schön, Adrian Wills and Brett Ninness. **System Identification of Nonlinear State-Space Models**. *Automatica*, 47(1):39-49, January 2011.



Consider a Bayesian SSM ( $\theta$  is now a random variable with a prior density  $p(\theta)$ )

$$\begin{aligned}x_{t+1} | x_t &\sim f_{\theta,t}(x_{t+1} | x_t), \\y_t | x_t &\sim g_{\theta,t}(y_t | x_t), \\x_1 &\sim \mu_{\theta}(x_1), \\ \theta &\sim p(\theta).\end{aligned}$$

**Identification problem:** Compute the posterior  $p(\theta, x_{1:T} | y_{1:T})$ , or one of its marginals.

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The **key challenge** is that there is no closed form expression available for the posterior.





Markov chain Monte Carlo (MCMC) methods allow us to generate samples from a target distribution by simulating a Markov chain.

**Gibbs sampling** (blocked) for SSMs amounts to iterating

- Draw  $\theta[m] \sim p(\theta \mid x_{1:T}[m-1], y_{1:T})$ ,
- Draw  $x_{1:T}[m] \sim p(x_{1:T} \mid \theta[m], y_{1:T})$ .

The above procedure results in a Markov chain,

$$\{\theta[m], x_{1:T}[m]\}_{m \geq 1}$$

with  $p(\theta, x_{1:T} \mid y_T)$  as its stationary distribution!



What would a Gibbs sampler for a general nonlinear/non-Gaussian SSM look like?

- Draw  $\theta[m] \sim p(\theta \mid x_{1:T}[m-1], y_{1:T})$ ;
- Draw  $x_{1:T}[m] \sim p(x_{1:T} \mid \theta[m], y_{1:T})$ .

OK!

Hard!

**Problem:**  $p(x_{1:T} \mid \theta[m], y_{1:T})$  not available!

**Idea:** Approximate  $p(x_{1:T} \mid \theta[m], y_{1:T})$  using a sequential Monte Carlo method!



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The particle filter provides an approximation of the filter PDF  $p(x_t | y_{1:t})$ , when the state evolves according to an SSM,

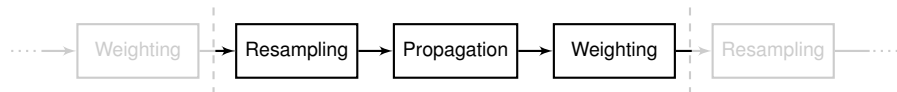
$$\begin{aligned}x_{t+1} | x_t &\sim f_t(x_{t+1} | x_t), \\y_t | x_t &\sim g_t(y_t | x_t), \\x_1 &\sim \mu(x_1).\end{aligned}$$

The particle filter maintains an empirical distribution made up  $N$  samples (particles) and corresponding weights

$$\hat{p}^N(x_t | y_{1:t}) = \sum_{i=1}^N w_t^i \delta_{x_t^i}(x_t).$$

*“Think of each particle as one simulation of the system state. Only keep the good ones.”*



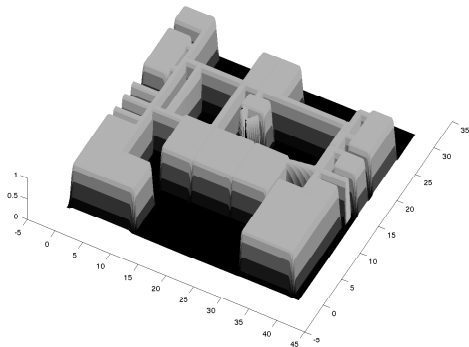


1. **Resampling:**  $\{x_{1:t-1}^i, w_{t-1}^i\}_{i=1}^N \rightarrow \{\tilde{x}_{1:t-1}^i, 1/N\}_{i=1}^N$ .
2. **Propagation:**  $x_t^i \sim R_t(x_t | \tilde{x}_{1:t-1}^i)$  and  $x_{1:t}^i = \{\tilde{x}_{1:t-1}^i, x_t^i\}$ .
3. **Weighting:**  $w_t^i = W_t(x_{1:t}^i)$ .

The result is a new weighted set of particles  $\{x_{1:t}^i, w_t^i\}_{i=1}^N$ .

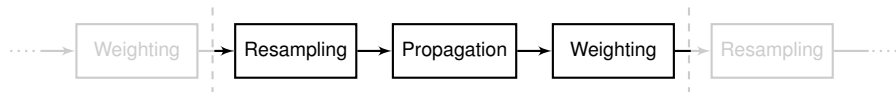


**Aim:** Compute the position of a person moving around indoors using sensors (inertial, magnetometer and radio) located in an ID badge and a map.



**Show movie**





## 1. Resampling + Propagation:

$$(a_t^i, x_t^i) \sim M_t(a_t, x_t) = \frac{w_{t-1}^{a_t}}{\sum_l w_{t-1}^l} R_t(x_t \mid x_{1:t-1}^a).$$

## 2. Weighting: $w_t^i = W_t(x_{1:t}^i)$ .

The result is a new weighted set of particles  $\{x_{1:t}^i, w_t^i\}_{i=1}^N$ .

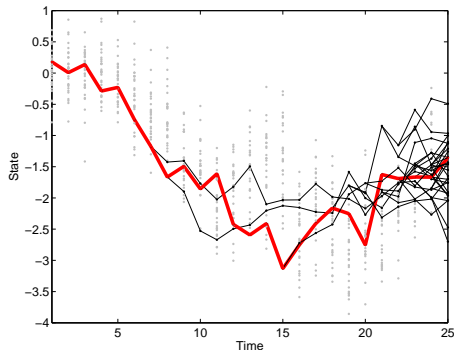


# The particle filter – illustrating particle degeneracy 16(32)





With  $P(x'_{1:T} = x^i_{1:T}) \propto w_T^i$  we get,  $x'_{1:T} \stackrel{\text{approx.}}{\sim} p(x_{1:T} | \theta, y_{1:T})$ .



Problems with this approach,

- Based on a PF  $\Rightarrow$  approximate sample.
- Does not leave  $p(\theta, x_{1:T} \mid y_{1:T})$  invariant!
- Relies on large  $N$  to be successful.
- A lot of wasted computations.

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To get around these problems,

Use a conditional particle filter (CPF). One prespecified path is retained throughout the sampler.

Christophe Andrieu, Arnaud Doucet and Roman Holenstein, **Particle Markov chain Monte Carlo methods**, *Journal of the Royal Statistical Society: Series B*, 72:269-342, 2010.



The idea underlying **PMCMC** is to make use of a certain SMC sampler to construct a Markov kernel leaving the joint smoothing distribution  $p(x_{1:T} \mid \theta, y_{1:T})$  invariant.

This Markov kernel is then used in a standard MCMC algorithm (e.g. Gibbs, results in the **Particle Gibbs (PG)**).

Three SMC samplers leaving  $p(x_{1:T} \mid \theta, y_{1:T})$  invariant:

1. **Conditional particle filter (CPF)**

Christophe Andrieu, Arnaud Doucet and Roman Holenstein, **Particle Markov chain Monte Carlo methods**, *Journal of the Royal Statistical Society: Series B*, 72:269-342, 2010.

2. **CPF with backward simulation (CPF-BS)**

Fredrik Lindsten and Thomas B. Schn. **On the use of backward simulation in the particle Gibbs sampler**. *Proc. of the 37th Internat. Conf. on Acoustics, Speech, and Signal Processing (ICASSP)*, Kyoto, Japan, March 2012.

3. **CPF with ancestor sampling (CPF-AS)**

Fredrik Lindsten, Michael I. Jordan and Thomas B. Schön, **Ancestor sampling for particle Gibbs**, *Proceedings of Neural Information Processing Systems (NIPS)*, Lake Tahoe, NV, US, December, 2012.



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**Algorithm** CPF w. ancestor sampling (CPF-AS), conditioned on  $x_{1:T}^*$

---

1. **Initialize** ( $t = 1$ ):

(a) Draw  $x_1^i \sim R_1^\theta(x_1)$  for  $i \neq N$  and set  $x_1^N = x_1^*$ .

(b) Set  $w_1^i = W_1^\theta(x_1^i)$  for  $i = 1, \dots, N$ .

2. **for**  $t = 2, \dots, T$ :

(a) Draw  $(a_t^i, x_t^i) \sim M_t^\theta(a_t, x_t)$  for  $i \neq N$  and set  $x_t^N = x_t^*$ .

(b) Draw  $a_t^N$  with  $P(a_t^N = i) \propto w_{t-1}^i p(x_t^* | \theta, x_{t-1}^i)$ .

(c) Set  $x_{1:t}^i = \{x_{1:t-1}^i, x_t^i\}$  and  $w_t^i = W_t^\theta(x_{1:t}^i)$  for  $i = 1, \dots, N$ .

---

(The red text highlights the difference to the standard PF)



## Theorem

For any  $N \geq 2$ , the procedure;

(i) Run CPF-AS( $x_{1:T}^*$ );

(ii) Sample  $P(x'_{1:T} = x^i_{1:T}) \propto w_T^i$ ;

defines a Markov kernel on  $X^T$  which leaves  $p(x_{1:T} \mid \theta, y_{1:T})$  invariant.

Three additional reasons for using CPF-AS:

1. Significantly improves the mixing compared to CPF.
2. The computational complexity is linear in  $N$ .
3. Opens up for non-Markovian models.



CPF

CPF-AS



**Bayesian identification:** Gibbs + CPF-AS = PG-AS

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## Algorithm PG-AS: Particle Gibbs with ancestor sampling

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1. **Initialize:** Set  $\{\theta[0], x_{1:T}[0]\}$  arbitrarily.
  2. **For**  $m \geq 1$ , **iterate:**
    - (a) Draw  $\theta[m] \sim p(\theta \mid x_{1:T}[m-1], y_{1:T})$ .
    - (b) Run CPF-AS( $x_{1:T}[m-1]$ ), targeting  $p(x_{1:T} \mid \theta[m], y_{1:T})$ .
    - (c) Sample with  $P(x_{1:T}[m] = x_{1:T}^i) \propto w_T^i$ .
- 

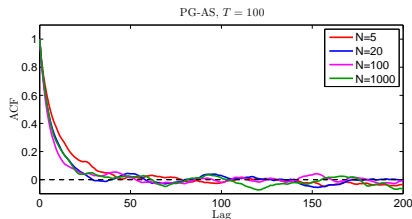
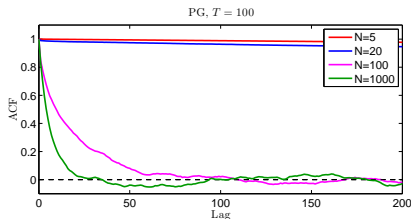
For any number of particles  $N \geq 2$ , the Markov chain  $\{\theta[m], x_{1:T}[m]\}_{m \geq 1}$  has stationary distribution  $p(\theta, x_{1:T} \mid y_{1:T})$ .



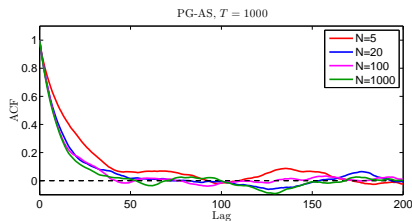
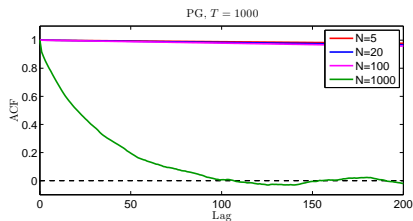
Consider the stochastic volatility model,

$$\begin{aligned}
 x_{t+1} &= 0.9x_t + w_t, & w_t &\sim \mathcal{N}(0, \theta), \\
 y_t &= e_t \exp\left(\frac{1}{2}x_t\right), & e_t &\sim \mathcal{N}(0, 1).
 \end{aligned}$$

Let us study the ACF for the estimation error,  $\hat{\theta} - \mathbb{E}[\theta \mid y_{1:T}]$



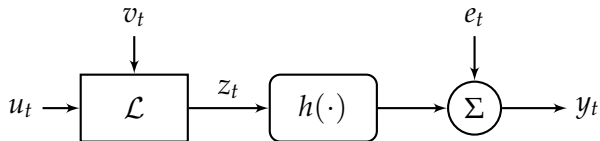




## Some observations:

- We want the ACF to decay to zero as rapidly as possible (indicates good mixing in the PG sampler).
- Note the superior mixing of PG-AS compared to PG-CPF (already for just  $N = 5$  particles!).





Parametric LGSS and a nonparametric static nonlinearity:

$$x_{t+1} = \underbrace{\begin{pmatrix} A & B \end{pmatrix}}_{\Gamma} \begin{pmatrix} x_t \\ u_t \end{pmatrix} + v_t, \quad v_t \sim \mathcal{N}(0, Q),$$

$$z_t = Cx_t.$$

$$y_t = h(z_t) + e_t,$$

$$e_t \sim \mathcal{N}(0, R).$$



Everything is learned from the data, by introducing the possibility to switch specific model components on and off.

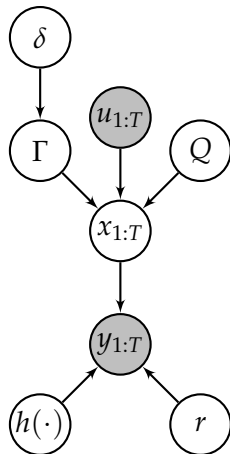
**“Parameters”**:  $\theta = \{A, B, Q, \delta, h(\cdot), r\}$ .

**Bayesian model** specified by priors

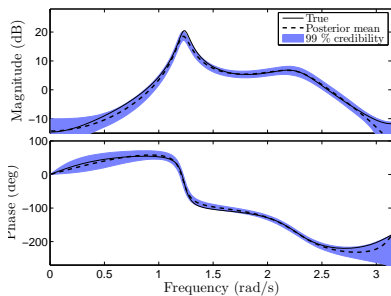
- Sparseness prior (ARD) on  $\Gamma = [A \ B]$ ,
- Inverse-Wishart prior on  $Q$  and  $r$
- Gaussian process prior on  $h(\cdot)$ ,

$$h(\cdot) \sim \mathcal{GP}(z, k(z, z')).$$

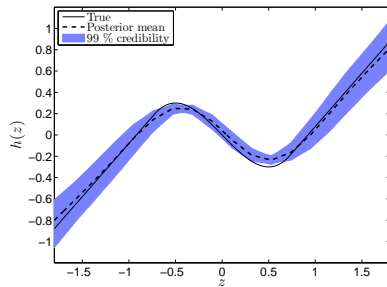
**Inference** using PG-AS with  $N = 15$  particles.  
 $T = 1\,000$  measurements. We ran 15\,000 MCMC iterations and discarded 5\,000 as burn-in.



Show movie



Bode diagram of the 4th-order linear system. Estimated mean (dashed black), true (solid black) and 99% credibility intervals (blue).



Static nonlinearity (non-monotonic), estimated mean (dashed black), true (black) and the 99% credibility intervals (blue).

Fredrik Lindsten, Thomas B. Schön and Michael I. Jordan. **Bayesian semiparametric Wiener system identification.** *Automatica*, 49(7): 2053-2063, July 2013.



Assume for the time being that we can sample from  $p_{\theta}(x_{1:T} | y_{1:T})$ .

**Stochastic approximation EM (SAEM):** Replace the E-step with,

$$\hat{Q}_m(\theta) = \hat{Q}_{m-1}(\theta) + \gamma_m \left( \frac{1}{M} \sum_{j=1}^M \log p_{\theta}(\tilde{x}_{1:T}^j, y_{1:T}) - \hat{Q}_{m-1}(\theta) \right),$$

where  $\tilde{x}_{1:T}^j \stackrel{\text{i.i.d.}}{\sim} p_{\theta}(x_{1:T} | y_{1:T})$  for  $j = 1, \dots, M$ .

SAEM converges to a maximum of  $p_{\theta}(y_{1:T})$  for any  $M \geq 1$  under standard stochastic approximation conditions.

B. Delyon, M. Lavielle and E. Moulines, **Convergence of a stochastic approximation version of the EM algorithm**, *The Annals of Statistics*, 27:94-128, 1999.



- **Bad news:** We cannot sample from  $p_{\theta}(x_{1:T} | y_{1:T})$ .
- **Good news:** It is enough to sample from a uniformly ergodic Markov kernel, leaving  $p_{\theta}(x_{1:T} | y_{1:T})$  invariant.

We can use CPF-AS to sample the states!

Results in an interesting and useful **combination** of frequentist and Bayesian ideas. We will see more combinations like this in the future.

Fredrik Lindsten. **An efficient stochastic approximation EM algorithm using conditional particle filters.** *Proceedings of the 38th International Conference on Acoustics, Speech, and Signal Processing (ICASSP)*, Vancouver, Canada, May 2013.



- EM-PS for ML learning in nonlinear SSMs.
- Conditional particle filters (CPF) are useful for identification!
- CPF-AS defines a kernel on  $X^T$  leaving  $p_{\theta}(x_{1:T} | y_{1:T})$  invariant.
- CPF-AS consists of two parts:
  - **Conditioning:** Ensures correct stationary distribution for any  $N$ .
  - **Ancestor sampling:** Mitigates path degeneracy and enables movement around the conditioned path.
- PG-AS for Bayesian inference and SAEM-AS for maximum likelihood inference. Both work with very few particles!

- 
- We are working on a book project,

Thomas B. Schön and Fredrik Lindsten, **Computational learning in dynamical systems**, 2013.

Send me an e-mail if you are interested in a draft.

- **Course:** [users.isy.liu.se/rt/schon/course\\_CIDS.html](http://users.isy.liu.se/rt/schon/course_CIDS.html)



## Forthcoming book

Thomas B. Schön and Fredrik Lindsten, **Computational learning in dynamical systems**, 2013.

## Novel introduction of PMCMC (very nice paper!)

Christophe Andrieu, Arnaud Doucet and Roman Holenstein, **Particle Markov chain Monte Carlo methods**, *Journal of the Royal Statistical Society: Series B*, 72:269-342, 2010.

## Self-contained introduction to BS and AS (not limited to SSMS)

Fredrik Lindsten and Thomas B. Schön, **Backward simulation methods for Monte Carlo statistical inference**, *Foundations and Trends in Machine Learning*, 2013 (accepted for publication).

## PG-AS (and the Wiener identification example)

Fredrik Lindsten, Michael I. Jordan and Thomas B. Schön, **Ancestor sampling for particle Gibbs**, *Proceedings of Neural Information Processing Systems (NIPS)*, Lake Tahoe, NV, US, December, 2012.

Fredrik Lindsten, Thomas B. Schön and Michael I. Jordan. **Bayesian semiparametric Wiener system identification**. *Automatica*, 2013, 49(): 2053-2063.

## ML identification of nonlinear SSMS (and Wiener example)

Thomas B. Schön, Adrian Wills and Brett Ninness. **System Identification of Nonlinear State-Space Models**. *Automatica*, 47(1):39-49, January 2011.

Adrian Wills, Thomas B. Schön, Lennart Ljung and Brett Ninness. **Identification of Hammerstein-Wiener Models**. *Automatica*, 49(1): 70-81, January 2013.

MATLAB code is available from our web-site.

